

# Parameter Estimation in Gaussian Process Regression for Deterministic Functions

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Introduction

Maximum likelihood estimation for deterministic  $f$

What if  $f$  is a sample from a GP?

Two future results

# Modelling Deterministic Functions

- Let  $\Omega \subset \mathbb{R}^d$ .
- Let  $f: \Omega \rightarrow \mathbb{R}$  be a **deterministic** function.
- Suppose that  $f$  generates the **noiseless data**

$$\mathcal{D}_N = \{(x_1, f(x_1)), \dots, (x_N, f(x_N))\} \text{ at some } x_i \in \Omega.$$

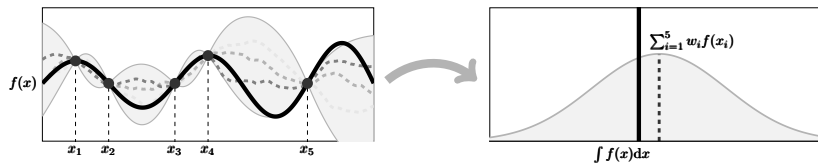
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In this talk we model  $f$  using a **Gaussian process**  $f_{\text{GP}}$ .

- $f$  need not be a sample from (or in any way related to)  $f_{\text{GP}}$ .
- The assumption that there is no noise is crucial.

# Motivation

- *Probabilistic numerics.* Provide quantification of epistemic uncertainty arising from discretisation in numerical approximation.
- *Modelling of computer experiments.* Predict the output of a complex and computationally expensive piece of code.
- *Bayesian optimisation.* Construct surrogates to an objective function.



- (K2020) **Karvonen, Wynne, Tronarp, Oates & Särkkä (2020)**. Maximum likelihood estimation and uncertainty quantification for Gaussian process approximation of deterministic functions. *SIAM/ASA Journal on Uncertainty Quantification*, 8(3):926–958.
  
- (K2021a) **Karvonen (2021)**. Small sample spaces for Gaussian processes. *arXiv:2103.03169*.
  
- (K2021b) **Karvonen (2021)**. Estimation of the scale parameter for a misspecified Gaussian process model. *arXiv:2110.02810*.
  
- + Two papers in preparation.

# Gaussian process interpolation I

- Let  $K: \Omega \times \Omega \rightarrow \mathbb{R}$  be a **positive-definite covariance kernel**.
- Model  $f$  as a **Gaussian process**  $f_{\text{GP}} \sim \text{GP}(0, K)$ .
- Condition  $f_{\text{GP}}$  on the **data**  $\mathcal{D}_N = \{(x_i, f(x_i))\}_{i=1}^N$ .

The resulting conditional GP,  $f_{\text{GP}} \mid \mathcal{D}_N$ , has the **mean**

$$s_{f,N}(x) := \mathbb{E}[f_{\text{GP}}(x) \mid \mathcal{D}_N] = f_N^\top K_N^{-1} k_N(x) \quad (1)$$

and **variance**

$$P_N(x)^2 := \text{Var}[f_{\text{GP}}(x) \mid \mathcal{D}_N] = K(x, x) - k_N(x)^\top K_N^{-1} k_N(x), \quad (2)$$

where

$$(f_N)_i = f(x_i), \quad (k_N(x))_i = K(x, x_i) \quad \text{and} \quad (K_N)_{ij} = K(x_i, x_j).$$

## Gaussian Process Interpolation II

- The conditional mean **interpolates** the data:

$$s_{f,N}(x_i) = f(x_i) \quad \text{for every } i = 1, \dots, N.$$

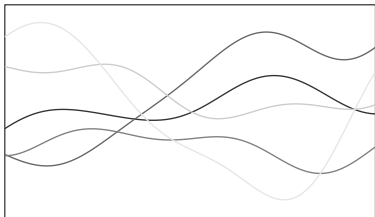
- The conditional variance **vanishes** at the data locations:

$$P_N(x_i)^2 = 0 \quad \text{for every } i = 1, \dots, N.$$

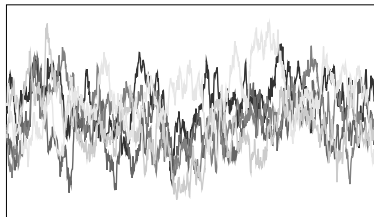
- The conditional variance **does not depend** on the data values  $f(x_i)$ .
- Properties of the kernel  $K$  define the properties of  $f_{\text{GP}}$ .

# Gaussian process priors

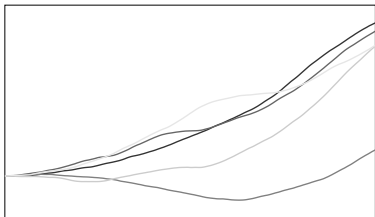
Gaussian:  $K(x, y) = e^{-(x-y)^2/2}$



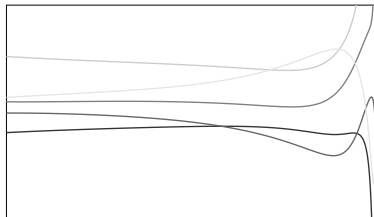
Matérn:  $K(x, y) = e^{-|x-y|}$



BM:  $K(x, y) = \frac{\min\{x,y\}^3}{3} + \frac{|x-y|\min\{x,y\}^2}{2}$

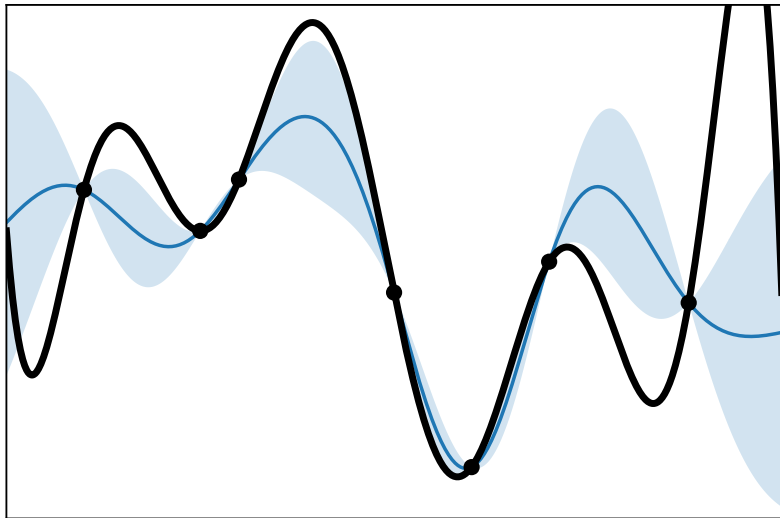


Hardy:  $K(x, y) = \frac{1}{1-xy}$





# Gaussian Process Conditional



# Objective

- Suppose that the prior covariance is parametric:  $f_{\text{GP}} \sim \text{GP}(0, K_{\theta})$ .
- The conditional process is

$$f_{\text{GP}} \mid \mathcal{D}_N \sim \text{GP}(s_{\theta, f, N}, P_{\theta, N}^2).$$

- For any  $a \in (0, 1)$ ,

$$\mathbb{P}\left[|f_{\text{GP}}(x) - s_{\theta, f, N}(x)| \leq c(a)P_{\theta, N}(x, x) \mid \mathcal{D}_N\right] = 1 - a.$$

- Compute hyperparameter estimates  $\theta(f, N)$  of  $\theta$ .

**Objective:** Understand the behaviour of (i)  $\theta(f, N)$  and (ii) the standard score

$$\frac{|f(x) - s_{\theta(f, N), f, N}(x)|}{P_{\theta(f, N), N}(x, x)} \quad \text{as } N \rightarrow \infty \quad (3)$$

for different  $f$  and hyperparameter estimation methods.

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# Reproducing kernel Hilbert spaces

## Reproducing kernel Hilbert space

Every covariance kernel  $K : \Omega \times \Omega \rightarrow \mathbb{R}$  induces a unique *reproducing kernel Hilbert space* (RKHS)  $\mathcal{H}(K)$  with inner product  $\langle \cdot, \cdot \rangle_K$ . The RKHS consists of functions  $g : \Omega \rightarrow \mathbb{R}$  and the kernel has the *reproducing property*

$$\langle g, K(\cdot, x) \rangle_K = g(x) \quad \text{for any } g \in \mathcal{H}(K) \text{ and } x \in \Omega.$$

The GP conditional moments are related to optimal interpolation in  $\mathcal{H}(K)$ .

## $s_{f,N}$ = minimum-norm interpolant

$$s_{f,N} = \arg \min_{s \in \mathcal{H}(K)} \{ \|s\|_K : s(x_i) = f(x_i) \text{ for every } i = 1, \dots, N \}$$

## $P_N(x, x) =$ worst-case error

$$P_N(x, x) = \sup_{\|g\|_K \leq 1} |g(x) - s_{g,N}(x)|$$

# Sobolev spaces

The **Sobolev space**  $H^\alpha(\mathbb{R}^d)$  consists of functions  $g \in L^2(\mathbb{R}^d)$  such that

$$\|g\|_\alpha^2 := \int_{\mathbb{R}^d} (1 + \|\xi\|^2)^\alpha |\widehat{g}(\xi)|^2 d\xi < \infty.$$

- For  $\Omega \subset \mathbb{R}^d$  the space  $H^\alpha(\Omega)$  is defined via restrictions.
- If  $\alpha > n + d/2$  for  $n \in \mathbb{N}$ , then  $H^\alpha(\mathbb{R}^d) \subset C^n(\mathbb{R}^d)$ .

## Sobolev kernel

Let  $\alpha > d/2$  and  $\Omega \subset \mathbb{R}^d$ . A kernel  $K : \Omega \times \Omega \rightarrow \mathbb{R}$  is a *Sobolev kernel of order  $\alpha$*  if its RKHS  $\mathcal{H}(K)$  is norm-equivalent ( $\simeq$ ) to  $H^\alpha(\Omega)$ .

A **Matérn kernel**

$$K(x, y) = \sigma^2 \frac{2^{1-\nu}}{\Gamma(\nu)} \left( \frac{\sqrt{2\nu} \|x - y\|}{\lambda} \right)^\nu \mathbf{K}_\nu \left( \frac{\sqrt{2\nu} \|x - y\|}{\lambda} \right)$$

of smoothness  $\nu > 0$  is a Sobolev kernel of order  $\nu + d/2$ .

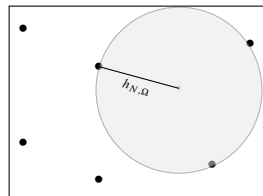
# Sobolev rates

Suppose that

- $\Omega \subset \mathbb{R}^d$  is bounded and sufficiently regular (e.g.,  $\Omega = [0, 1]^d$ ).
- The points  $x_1, \dots, x_N$  are **quasi-uniform**: the *fill-distance*

$$h_{N,\Omega} = \sup_{x \in \Omega} \min_{i=1, \dots, N} \|x - x_i\|$$

is of order  $N^{-1/d}$  as  $N \rightarrow \infty$ .



## Theorem (from approximation theory)

Let  $d/2 < \beta \leq \alpha$ . Suppose that  $\mathcal{H}(K) \simeq H^\alpha(\Omega)$  and  $f \in H^\beta(\Omega)$ . Then

$$\sup_{x \in \Omega} |f(x) - s_{f,N}(x)| \leq C_1 \|f\|_\beta N^{-\beta/d+1/2} \quad (4)$$

and

$$C_2 N^{-\alpha/d+1/2} \leq P_N(x) \leq C_3 N^{-\alpha/d+1/2} \quad \text{for } x \notin \{x_i\}_{i=1}^\infty. \quad (5)$$

## Why parameter estimation is necessary

If  $K$  is a Sobolev kernel of order  $\alpha$  and  $f \in H^\beta(\Omega)$  for  $\beta \leq \alpha$ , then

$$\frac{|f(x) - s_{f,N}(x)|}{P_N(x,x)} \leq \frac{C_1 \|f\|_\beta N^{-\beta/d+1/2}}{C_2 N^{-\alpha/d+1/2}} = C_4 N^{(\alpha-\beta)/d}.$$

$\implies$  If  $f \notin \mathcal{H}(K) \simeq H^\alpha(\Omega)$ , it may be necessary estimate kernel parameters in order to prevent overconfidence:

$$\frac{\alpha - \beta}{d} > 0.$$

## Maximum likelihood estimation

The log-likelihood function is

$$\ell(\theta) = -\frac{1}{2} \left[ f_N^\top K_{\theta,N}^{-1} f + \log \det K_{\theta,N} + N \log(2\pi) \right],$$

where  $(f_N)_i = f(x_i)$  and  $(K_{\theta,N})_{ij} = K_\theta(x_i, x_j)$ .

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Fix  $K$  and use the simple parametrisation  $K_\sigma = \sigma^2 K$  for  $\sigma \geq 0$ . Then

$$s_{\sigma,f,N}(x) = s_{f,N}(x) \quad \text{and} \quad P_{\sigma,N}(x, y) = \sigma P_N(x, y).$$

$$\sigma_{\text{ML}}(f, N) = \arg \max_{\sigma \geq 0} \ell(\sigma) = \sqrt{\frac{f_N^\top K_N^{-1} f_N}{N}} = \frac{\|s_{f,N}\|_K}{\sqrt{N}} \quad (6)$$

and

$$\text{standard score} = \frac{|f(x) - s_{f,N}(x)|}{\sigma_{\text{ML}}(f, N) P_N(x, x)} \quad (7)$$



## First results — $f \in \mathcal{H}(K)$

### Proposition (Proposition 3.1 in K2020)

If  $f \in \mathcal{H}(K)$  and  $f(x_i) \neq 0$  for some  $x_i$ , then there is  $c > 0$  such that

$$c N^{-1/2} \leq \sigma_{\text{ML}}(f, N) \leq \|f\|_K N^{-1/2}. \quad (8)$$

### Theorem (Theorem 3.2 in K2020)

If  $f \in \mathcal{H}(K)$  and  $f(x_i) \neq 0$  for some  $x_i$ , then there is  $C > 0$  such that

$$\frac{|f(x) - s_{f,N}(x)|}{\sigma_{\text{ML}}(f, N) P_N(x, x)} \leq C \sqrt{N}. \quad (9)$$

At most “slow” (i.e.,  $\sqrt{N}$ ) overconfidence if  $f \in \mathcal{H}(K)$ .

## Second result — $f \in H^\beta(\Omega)$ and $\mathcal{H}(K) \simeq H^\alpha(\Omega)$

Let  $\beta \in (d/2, \alpha]$ . Suppose that  $\Omega$  is regular and  $\{x_i\}_{i=1}^\infty$  are quasi-uniform.

### Proposition (Proposition 4.5 in K2020)

If  $\mathcal{H}(K) \simeq H^\alpha(\Omega)$  and  $f \in H^\beta(\Omega)$ , then

$$\sigma_{\text{ML}}(f, N) \leq C_1 N^{(\alpha-\beta)/d-1/2} \|f\|_{H^\beta(\Omega)}. \quad (10)$$

$\implies$  the behaviour of  $\sigma_{\text{ML}}(f, N)$  tells how “far” from  $\mathcal{H}(K)$  the function  $f$  is.

### Theorem (Theorem 4.10 in K2020)

If  $\mathcal{H}(K) \simeq H^\alpha(\Omega)$  and  $f$  has “exact smoothness”  $\beta$  for  $\lfloor \beta \rfloor > d/2$ , then

$$\frac{|f(x) - s_{f,N}(x)|}{\sigma_{\text{ML}}(f, N) P_N(x, x)} \leq C_2(f) (\log N)^{\alpha/(2\beta)} \sqrt{N}. \quad (11)$$

At most “slow” (i.e.,  $\approx \sqrt{N}$ ) overconfidence if  $f$  is rougher than  $\mathcal{H}(K)$ .

## Some implications

- Overconfidence cannot be ruled out, but at least it cannot be overly severe: the standard score is approximately  $O(N^{1/2})$ .
- Simple maximum likelihood estimation of a scaling parameter provides strong protection against smoothness misspecification.
- MLE does not detect undersmoothing by the model:

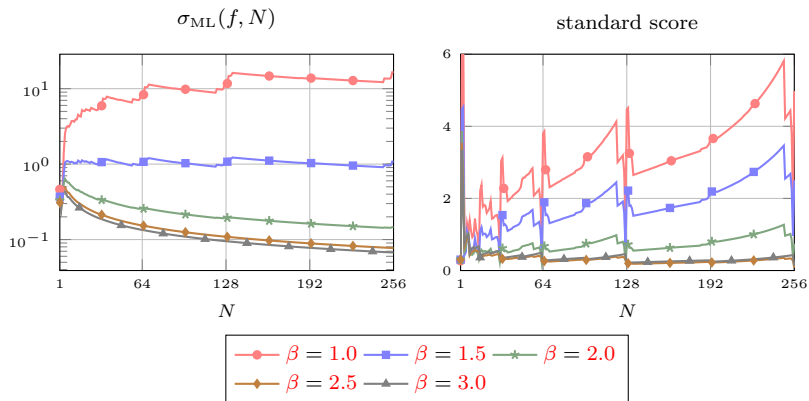
$$\sigma_{\text{ML}}(f, N) \asymp N^{-1/2} \quad \text{for every non-zero } f \in \mathcal{H}(K).$$

- It can be shown (**Theorem 4.11 in K2020**) that underconfidence occurs if there is (roughly speaking) sufficient undersmoothing:

$$\mathcal{H}(K) \simeq H^\alpha(\Omega) \quad \text{and} \quad f \in H^{2\alpha}(\Omega).$$

# Numerical results: confidence intervals

$\mathcal{H}(K) \simeq H^2([0, 1])$  and  $f \in H^\beta([0, 1])$  for  $\beta = \frac{2}{2}, \frac{3}{2}, \dots, \frac{6}{2}$ .



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# Sample path properties

## Theorem (e.g., Steinwart 2019)

Let  $f_{\text{GP}} \sim \text{GP}(0, K)$  and suppose that  $K$  is a Sobolev kernel of order  $\alpha > d/2$ . Then

$$\mathbb{P}[f_{\text{GP}} \in H^{\beta}(\Omega)] = 0 \quad \text{if} \quad \beta \geq \alpha - d/2$$

and

$$\mathbb{P}[f_{\text{GP}} \in H^{\beta}(\Omega)] = 1 \quad \text{if} \quad \beta < \alpha - d/2.$$

The samples of the Gaussian process  $f_{\text{GP}}$  are therefore “ $d/2$  less smooth” than the RKHS  $\mathcal{H}(K)$ .

⇒ Samples are *not* in the RKHS  $\mathcal{H}(K) \simeq H^{\alpha}(\Omega)$ !

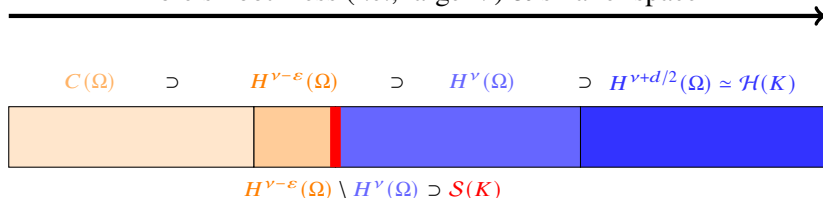
⇒ Samples have “exact smoothness”  $\alpha - d/2$  if  $\mathcal{H}(K) \simeq H^{\alpha}(\Omega)$ .

# Matérn kernels

Let  $K$  be a Matérn kernel of smoothness  $\nu > 0$  and  $\Omega \subset \mathbb{R}^d$  sufficiently regular. Then

- $\mathcal{H}(K)$  is norm-equivalent to the Sobolev space  $H^{\nu+d/2}(\Omega)$ .
- $\mathbb{P}[f_{\text{GP}} \in H^{\nu-\varepsilon}(\Omega)] = 1$  if and only if  $\varepsilon > 0$ .

more smoothness (i.e., larger  $\nu$ ) & smaller space



The samples of  $f_{\text{GP}}$  can be thought of as elements of the “Sobolev slice”

$$\mathcal{S}(K) \subset H^{\nu-\varepsilon}(\Omega) \setminus H^\nu(\Omega) \quad \text{for every } \varepsilon > 0.$$

## Expected scale MLE for samples

Suppose that the data-generating function is a Gaussian process

$$f = f_{\text{GP}}^* \sim \text{GP}(0, R)$$

but the model is  $f_{\text{GP}} \sim \text{GP}(0, K)$ . Then

$$\mathbb{E}_{f_{\text{GP}}^*} [\sigma_{\text{ML}}(f_{\text{GP}}^*, N)^2] = \mathbb{E}_{f_{\text{GP}}^*} \left[ \frac{(f_{\text{GP},N}^*)^\top K_N^{-1} f_{\text{GP},N}^*}{N} \right] = \frac{\text{trace}(R_N K_N^{-1})}{N}.$$

### Theorem (Theorem 4.2 in K2021b)

Suppose that  $\mathcal{H}(K) \simeq H^\alpha([0, 1]^d)$  and  $\mathcal{H}(R) \simeq H^{\alpha_0}([0, 1]^d)$  for  $\alpha \geq \alpha_0 > d/2$ . If the points  $\{x_i\}_{i=1}^\infty$  are quasi-uniform, then

$$C_1 N^{2(\alpha - \alpha_0)/d} \leq \mathbb{E}_{f_{\text{GP}}^*} [\sigma_{\text{ML}}(f_{\text{GP}}^*, N)^2] \leq C_2 N^{2(\alpha - \alpha_0)/d}.$$



## Comparison to the deterministic case

Suppose that  $\mathcal{H}(K) \simeq H^\alpha([0, 1]^d)$  and that  $\{x_i\}_{i=1}^\infty$  are quasi-uniform.

Recall that the samples “have smoothness  $\alpha - d/2$ ”.

### Deterministic

Let  $f \in H^{\alpha-d/2}([0, 1]^d)$  so that  $f$  does not have less smoothness than the samples of  $f_{\text{GP}} \sim \text{GP}(0, K)$ . For  $\beta = \alpha - d/2$  we get

$$\sigma_{\text{ML}}(f, N)^2 \leq C_1 N^{2(\alpha-\beta)/d-1} \|f\|_{H^\beta(\Omega)}^2 = C_1 \|f\|_{H^{\alpha-d/2}(\Omega)}^2. \quad (12)$$

### Stochastic

Let  $f = f_{\text{GP}}^* \sim \text{GP}(0, R)$  for  $R$  such that  $\mathcal{H}(R) \simeq H^\alpha([0, 1]^d)$ . Then  $f_{\text{GP}} \sim \text{GP}(0, K)$  and  $f_{\text{GP}}^*$  have similar paths. For  $\alpha_0 = \alpha$  we get

$$\mathbb{E}_{f_{\text{GP}}^*} [\sigma_{\text{ML}}(f_{\text{GP}}^*, N)^2] \asymp N^{(\alpha-\alpha_0)/d} = 1. \quad (13)$$

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# Is cross-validation better than MLE?

The **leave-one-out cross-validated** estimate of  $\sigma$  is

$$\sigma_{\text{cv}}(f, N) = \frac{1}{N} \sum_{i=1}^N \left( \frac{f(x_i) - s_{f, N \setminus i}(x_i)}{P_{N \setminus i}(x_i, x_i)} \right)^2,$$

where  $s_{f, N \setminus i}$  and  $P_{N \setminus i}$  are the GP conditional mean and std based on the data set  $\mathcal{D}_N \setminus \{(x_i, f(x_i))\}$ .

- Recall that for any non-zero  $f \in \mathcal{H}(K)$  we have

$$\sigma_{\text{ML}}(f, N) \asymp N^{-1/2} \quad \text{for any non-zero } f \in \mathcal{H}(K).$$

- At least in some cases it can be proved<sup>1</sup> that the rate of decay of

$$\sigma_{\text{cv}}(f, N) \quad \text{depends on } f \in \mathcal{H}(K).$$

$\implies$  Cross-validation is **more sensitive** to the smoothness of  $f$  than MLE.

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<sup>1</sup>Work in progress with M. Naslidnyk, M. Mahsereci and M. Kanagawa.

## Estimating the lengthscale parameter

The kernel  $K$  is stationary if it can be written as

$$K(x, y) = \Phi\left(\frac{x - y}{\lambda}\right) \quad \text{for some} \quad \Phi: \mathbb{R}^d \rightarrow \mathbb{R},$$

where  $\lambda > 0$  is the **lengthscale** parameter.

### **Theorem (to appear in a paper with C. Oates)**

Let  $N \geq 2$  and suppose that  $\mathcal{H}(K) \simeq H^\alpha(\mathbb{R}^d)$  for some  $\alpha > d/2$ . If the data vector is **constant**,

$$f_N = (c, \dots, c) \in \mathbb{R}^N \quad \text{for some} \quad c \in \mathbb{R},$$

then

$$\lambda_{\text{ML}} = \infty.$$

# Conclusion

- Simple MLE of the scaling parameter  $\sigma$  in  $K_\sigma(x, y) = \sigma^2 K(x, y)$  provides **protection against misspecification**.
- **Overconfidence is possible** but at least it cannot be **too severe**.
- Samples from a GP are elements of a “small” set of functions. This is manifested in the sample results being “nicer”.
- Cross-validation may be **more sensitive** to the smoothness of  $f$  than MLE.

## Additional references

- **Lukić & Beder (2001)**. Stochastic processes with sample paths in reproducing kernel Hilbert spaces. *Transactions of the American Mathematical Society*, 353(10):3945–3969.
- **Narcowich, Ward & Wendland (2006)**. Sobolev error estimates and a Bernstein inequality for scattered data interpolation via radial basis functions. *Constructive Approximation*, 24(2):175–186.
- **Bachoc (2013)**. Cross validation and maximum likelihood estimations of hyper-parameters of Gaussian processes with model misspecification. *Computational Statistics & Data Analysis*, 66:55–69.
- **Xu & Stein (2017)**. Maximum likelihood estimation for a smooth Gaussian random field model. *SIAM/ASA Journal on Uncertainty Quantification*, 5(1):138–175.
- **Steinwart (2019)**. Convergence types and rates in generic Karhunen-Loève expansions with applications to sample path properties. *Potential Analysis*, 51(3):361–395.
- **Hadji & Szábo (2019)**. Can we trust Bayesian uncertainty quantification from Gaussian process priors with squared exponential covariance kernel? *arXiv:2002.01381*.
- **Teckentrup (2020)**. Convergence of Gaussian process regression with estimated hyper-parameters and applications in Bayesian inverse problems. *SIAM/ASA Journal on Uncertainty Quantification*, 8(4):1310–1337.
- **Wang (2020)**. On the inference of applying Gaussian process modeling to a deterministic function. *arXiv:2002.01381*.

Thank you for your attention!